

# Refinements of Berry-Esseen inequality in terms of Lyapunov coefficients

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## Setting

$X_1, \dots, X_n$  independent random variables,  $\mathbb{E}X_k = 0$ ,  $\text{Var}(X_k) = \sigma_k^2$ ,

$$S_n = X_1 + \dots + X_n.$$

**Normalization:**  $\mathbb{E}S_n^2 = \sigma_1^2 + \dots + \sigma_n^2 = 1$ .

**Distribution functions:**  $F_n(x) = \mathbb{P}\{S_n \leq x\}$ .

**Kolmogorov distance from the standard normal law:**

$$\Delta_n = \sup_x |F_n(x) - \Phi(x)|,$$

where

$$\Phi(x) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-y^2/2} dy, \quad x \in \mathbb{R}.$$

**IID model:**  $X_k = \xi_k / \sqrt{n}$  with iid  $\xi_k$  such that  $\mathbb{E}\xi_1 = 0$ ,  $\mathbb{E}\xi_1^2 = 1$ .

**Weighted sums iid:**  $X_k = a_k \xi_k$  ( $a_1^2 + \dots + a_n^2 = 1$ ,  $a_k \in \mathbb{R}$ ).

# CLT and Lyapunov coefficients

Lyapunov coefficients:

$$L_p = \sum_{k=1}^n \mathbb{E} |X|^p, \quad p > 2.$$

IID model:

$$L_p = n^{-\frac{p-2}{2}} \beta_p, \quad \beta_p = \mathbb{E} |\xi_1|^p.$$

**Theorem** (Berry-Esseen inequality). Up to some absolute constant  $c > 0$ ,

$$\Delta_n \leq cL_3.$$

In particular, in the iid model,

$$\Delta_n \leq \frac{c\beta_3}{\sqrt{n}}.$$

**Note:**  $L_3 \geq \frac{1}{\sqrt{n}}$ .

# Truncated Lyapunov coefficients

Definition:

$$R_p = \sum_{k=1}^n \mathbb{E} \min\{1, |X_k|^{p-2}\} X_k^2, \quad p > 2.$$

Note:  $\frac{c}{\sqrt{n}} \leq R_3 \leq \min\{1, L_3\}$ .

Theorem (improved Berry-Esseen inequality):  $\Delta_n \leq cR_3$ .

History in the typical model of increasing sums:  $S_n = \frac{1}{B_n} \sum_{k=1}^n \xi_k$

M.L. Katz (1963), V.V. Petrov (1965), Ju.P. Studnev (1958, 1965), L.V. Osipov (1965), W. Feller (1968)

Constants:  $c = 2.02$  (V.Yu. Korolev, S.V. Popov 2011),  $c = 1.87$  (A. Dorofeeva, V.Yu. Korolev 2017)

Reproved (as investigated by R.A. Gabdullin, V.A. Makarenko, I.G. Shevtsova 2018):

H. Paditz (1986), A.D. Barbour and P. Hall (2001 Stein method), L.H.Y. Chen and Q.M. Shao (2001)

# Connection with theorems by Lindeberg and Osipov

Consider the sums

$$S_n = \frac{1}{B_n} \sum_{k=1}^n \xi_k, \quad \mathbb{E}\xi_k = 0, \quad B_n^2 = \sum_{k=1}^n \mathbb{E}\xi_k^2, \quad V_k(x) = \mathbb{P}\{X_k \leq x\}.$$

For  $\varepsilon > 0$ , define

$$\Lambda_n(\varepsilon) = \frac{1}{B_n} \sum_{k=1}^n \int_{|x| > \varepsilon B_n} x^2 dV_k(x), \quad L_n(\varepsilon) = \frac{1}{B_n} \sum_{k=1}^n \int_{|x| \leq \varepsilon B_n} |x|^3 dV_k(x) \leq \varepsilon.$$

**Theorem (Lindeberg):** If  $\Lambda_n(\varepsilon) \rightarrow 0$  for any  $\varepsilon > 0$ , then  $\Delta_n \rightarrow 0$  as  $n \rightarrow \infty$ .

**Theorem (Osipov):**

$$|\Delta_n| \leq C (\Lambda_n(\varepsilon) + L_n(\varepsilon)).$$

**Notes:** Let  $X_k = \frac{1}{B_n} \xi_k$ . Gabdullin-Makarenko-Shevtsova 2018: For any  $\varepsilon > 0$ ,

$$R_3 = \Lambda_n(1) + L_n(1) \leq \Lambda_n(\varepsilon) + L_n(\varepsilon).$$

## Moments smaller than 3

**Moments**  $p = 2 + \delta$ ,  $0 < \delta < 1$ .

Since  $\min(1, |x|) x^2 \leq |x|^{2+\delta}$  for all  $x \in \mathbb{R}$ , we have

$$R_3 \leq L_{2+\delta}.$$

**Corollary:**

$$\Delta_n \leq cL_{2+\delta}.$$

In particular, in the iid model,

$$\Delta_n \leq \frac{c\beta_{2+\delta}}{n^{\delta/2}}.$$

**Notes:** The case  $0 < \delta < 1$ : Lyapunov (according to Osipov)  
**Necessary and sufficient conditions** for  $\Delta_n = O(n^{-\delta/2})$ : I. A. Ibragimov (1966).

# Optimality of Berry-Esseen bounds

Weighted sums:

$$S_n = a_1 \xi_1 + \dots + a_n \xi_n, \quad a_1^2 + \dots + a_n^2 = 1 \quad (a_k \in \mathbb{R})$$

with iid  $\xi_k$  such that  $\mathbb{E}\xi_1 = 0$ ,  $\mathbb{E}\xi_1^2 = 1$ . Put  $\alpha_3 = \mathbb{E}\xi_1^3$ ,  $\beta_p = \mathbb{E}|\xi_1|^p$ .

**Theorem 1.** a) Let  $\alpha_3 \neq 0$  and  $\beta_4 < \infty$ . If all  $a_k \geq 0$ , then with  $c' = c'(\alpha_3, \beta_4)$

$$c' L_3 \leq \Delta_n \leq c L_3.$$

b) If  $\beta_4 \neq 3$  and  $\beta_5 < \infty$ , then with  $c = c'(\beta_4, \beta_5)$

$$c' L_4 \leq \Delta_n \leq c L_3.$$

Note:

$$L_3 = \beta_3 \sum_{k=1}^n |a_k|^3 \geq \beta_3 (\max_k |a_k|)^3,$$

implying

$$L_4 \leq c L_3^{4/3}, \quad c = \beta_4 \beta_3^{-1/3}.$$

## New Berry-Esseen-type bounds

As before, let  $X_1, \dots, X_n$  be independent r.v. with mean zero, and let  $\mathbb{E}S_n^2 = 1$ .

**Theorem 2.** Suppose that  $X_k$  have finite 4-th moments with  $\mathbb{E}X_k^3 = 0$ . For any  $\delta \in (0, 1]$ ,

$$\Delta_n \leq c \left( \frac{1}{\delta} L_4 + L_{2+\delta}^{1/\delta} \right).$$

Moreover, if the distributions of  $X_k$  are symmetric and have finite absolute moments of order  $2 + \delta$ , then

$$\Delta_n \leq c \left( \frac{1}{\delta} R_4 + L_{2+\delta}^{1/\delta} \right).$$

**Notes.** If  $\delta = 1$ , we return to the Berry-Esseen theorem (symmetric case). In general,  $\delta \rightarrow L_{2+\delta}^{1/\delta}$  is non-decreasing. In particular,  $L_3 \leq \sqrt{L_4}$ .



## Examples

Let the i.i.d. random variables  $\xi_k$  have  $\mathbb{E}\xi_1 = \mathbb{E}\xi_1^3 = 0$ ,  $\mathbb{E}\xi_1^2 = 1$ ,  $\mathbb{E}\xi_1^4 < \infty$ . For  $0 < \alpha < \frac{1}{2}$ , consider

$$S_n = \frac{1}{b_n} \sum_{k=1}^n \frac{1}{k^\alpha} \xi_k, \quad b_n \sim n^{\frac{1}{2}-\alpha}.$$

**Main case:**  $\frac{1}{3} < \alpha < \frac{1}{2}$ . Choose  $\delta < \frac{1}{\alpha} - 2$ ). Then

$$L_3 \sim n^{-3(\frac{1}{2}-\alpha)}, \quad L_4 \sim n^{-4(\frac{1}{2}-\alpha)} \sim L_3^{4/3}, \quad L_{2+\delta}^{1/\delta} \sim \frac{1}{\sqrt{n}} = o(L_3).$$

Therefore, Th2 improves the usual B-E bound and yields

$$\Delta_n \leq c \left( L_4 + \frac{1}{\sqrt{n}} \right).$$

Moreover, by Th1,

$$\Delta_n \sim L_4, \quad \frac{3}{8} \leq \alpha < \frac{1}{2}.$$

**Note:** For  $\frac{1}{4} < \alpha < \frac{1}{3}$ ,

$$L_3 \sim L_{2+\delta}^{1/\delta} \sim \frac{1}{\sqrt{n}}, \quad L_4 = o(L_3).$$

## Chebyshev–Edgeworth correction

Let  $X_k$  have finite absolute moments of an integer order  $p \geq 4$ , and  $\Phi_{p-1}(x)$  denote the Chebyshev–Edgeworth correction for  $F_n(x)$  based on moments of  $X_k$  up to order  $p - 1$ .

First corrections.

$$\Phi_3(x) = \Phi(x) - \frac{\gamma_3}{3!} (x^2 - 1)\varphi(x), \quad \gamma_3 = \sum_{k=1}^n \mathbb{E}X_k^3,$$

$$\Phi_4(x) = \Phi(x) - \frac{\gamma_4}{4!} (x^3 - 3x)\varphi(x), \quad \gamma_4 = \sum_{k=1}^n (\mathbb{E}X_k^4 - 3) \quad (\text{if } \mathbb{E}X_k^3 = 0).$$

Theorem 3.

$$\sup_x |F_n(x) - \Phi_{p-1}(x)| \leq c_p \left( \delta^{-\frac{p-2}{2}} L_p + L_{2+\delta}^{1/\delta} \right), \quad 0 < \delta \leq 1.$$

Previous example:

$$\sup_x |F_n(x) - \Phi_{p-1}(x)| \leq c_{p,\alpha} \frac{\beta_p}{\sqrt{n}}, \quad p > \frac{1}{1 - 2\alpha}.$$

## Normal approximation for characteristic functions

Let  $X_1, \dots, X_n$  be independent r.v. with mean zero, and let  $\mathbb{E}S_n^2 = 1$ . Put  $f_n(t) = \mathbb{E} e^{itS_n}$ .

Esseen's smoothing:

$$c\Delta_n \leq \int_0^T \left| \frac{f_n(t) - e^{-t^2/2}}{t} \right| dt + \frac{1}{T}, \quad T > 0.$$

It is well-known that

$$|f_n(t) - e^{-t^2/2}| \leq cL_3 \min(1, |t|^3) e^{-t^2/6}, \quad |t| \leq \frac{1}{L_3}.$$

**Lemma 1** (Osipov). We have

$$|f_n(t) - e^{-t^2/2}| \leq cR_3 \min(1, |t|^3) e^{-t^2/6}, \quad |t| \leq \frac{1}{32R_3}.$$

**Lemma 2.** If the distributions of  $X_k$  are symmetric about the origin, then

$$|f_n(t) - e^{-t^2/2}| \leq cR_4 \min(1, |t|^3) e^{-t^2/6}, \quad |t| \leq \frac{1}{32R_3}.$$

## Characteristic functions for single random variables

Let  $X$  be a random variable with  $\mathbb{E}X = 0$ ,  $\text{Var}(X) = \sigma^2$ , and let

$$f(t) = \mathbb{E}e^{itX}, \quad t \in \mathbb{R}.$$

**Lemma 3.** For all  $t \in \mathbb{R}$ ,

$$f(t) = 1 - \frac{\sigma^2 t^2}{2} + \theta t^2 \mathbb{E} \min\{1, |tX|\} X^2, \quad |\theta| \leq 2.$$

Moreover, if the distribution of  $X$  is symmetric, then

$$f(t) = 1 - \frac{\sigma^2 t^2}{2} + \theta t^2 \mathbb{E} \min\{1, (tX)^2\} X^2, \quad |\theta| \leq 4.$$

## Cox-Kemperman inequality

**Lemma 4** (C-K 1983). Given independent random variables  $X$  and  $Y$  with mean zero and finite absolute moments of order  $p \geq 2$ ,

$$\mathbb{E} |X + Y|^p \leq 2^{p-2} (\mathbb{E} |X|^p + \mathbb{E} |Y|^p).$$

**Cases of equality:**  $X$  and  $Y$  have a symmetric Bernoulli distribution.

**Proof:** For all  $x, y \in \mathbb{R}$ ,

$$|x + y|^p \leq 2^{p-2} (|x|^p + |y|^p + x \operatorname{sign}(y) |y|^{p-1} + y \operatorname{sign}(x) |x|^{p-1}).$$

## Upper bounds for characteristic functions

**Lemma 5.** We have

$$|f_n(t)| \leq 2e^{-t^2/4}, \quad |t| \leq \frac{1}{32R_3}.$$

Proof: Apply Lemma 3.

**Lemma 6.** For all  $\delta \in (0, 2]$ ,

$$|f_n(t)| \leq e^{-\delta t^2/6}, \quad |t| \leq \frac{1}{L_{2+\delta}^{1/\delta}}.$$

In particular,

$$|f_n(t)| \leq e^{-t^2/6}, \quad |t| \leq \frac{1}{L_3}.$$

## Proof of Lemma 6

Let  $X_1, \dots, X_n$  be independent r.v. with mean zero, variances  $\sigma_k^2$  such that  $\mathbb{E}S_n^2 = 1$ . Put  $v_k(t) = \mathbb{E}e^{itX_k}$ , so that

$$f_n(t) = \mathbb{E}e^{itS_n} = v_1(t) \dots v_n(t).$$

We use

$$\cos x \leq 1 - \frac{1}{2}x^2 + \frac{1}{2} \cdot 12^{-\delta/2} |x|^{2+\delta}, \quad x \in \mathbb{R}.$$

Applying C-K inequality, this gives

$$\begin{aligned} |v_k(t)|^2 &= \mathbb{E} \cos(t(X_k - Y_k)) \\ &\leq 1 - \sigma_k^2 t^2 + \frac{1}{2} \cdot 12^{-\delta/2} |t|^{2+\delta} \mathbb{E} |X_k - Y_k|^{2+\delta} \\ &\leq 1 - \sigma_k^2 t^2 + 3^{-\delta/2} |t|^{2+\delta} \mathbb{E} |X_k|^{2+\delta} \\ &\leq \exp \left\{ -\sigma_k^2 t^2 + 3^{-\delta/2} |t|^{2+\delta} \mathbb{E} |X_k|^{2+\delta} \right\}. \end{aligned}$$

## Proof of Lemma 6 (continuation)

Thus,

$$|v_k(t)| \leq \exp \left\{ -\frac{\sigma_k^2 t^2}{2} + \frac{1}{2} \cdot 3^{-\delta/2} |t|^{2+\delta} \mathbb{E} |X_k|^{2+\delta} \right\}.$$

Multiplying over all  $k \leq n$ ,

$$|f_n(t)| \leq \exp \left\{ -\frac{t^2}{2} + \frac{1}{2} \cdot 3^{-\delta/2} |t|^{2+\delta} L_{2+\delta} \right\}.$$

As a result, if  $|t|^\delta L_{2+\delta} \leq 1$ ,

$$|f_n(t)| \leq \exp \left\{ -\frac{1}{2} (1 - 3^{-\delta/2}) t^2 \right\}.$$

To simplify, use  $1 - 3^{-\delta/2} \geq \frac{1}{3} \delta$  for  $0 < \delta \leq 2$ .